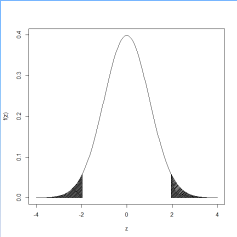


STATS 101
Introductory Statistics

Lecture 6
Confidence Interval Estimation
1-sample problems



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In this lecture, following will be discussed:

- Population of interest
- Independent random sample from the population
- Use SAMPLE data to draw conclusions about the POPULATION of interest

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STATISTICAL POPULATION

Look at a few examples:

1. A realtor needs to find the average selling price of homes in Las Vegas Metro.

POPULATION = {all single family homes
in Las Vegas
Metro}

VARIABLE of interest = price of home.

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2. A consumer group is interested in the number of failures of a new model of inline skates.

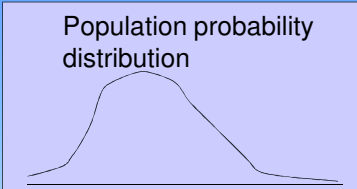
POPULATION = {all inline skates of this model}

VARIABLE of interest = # of failed inline skates in a given sample

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Population probability distribution



Sample = $\{x_1, x_2, \dots, x_n\}$

STATISTICAL INFERENCE : use sample information to draw conclusions about the population distribution.

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TWO APPROACHES

1. Confidence Interval Estimation
2. Testing Statistical Hypotheses

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Confidence Interval Estimation of Mean

Example 1: CI Estimation of the mean selling price of homes in a large city

Suppose the realtor collects a representative random sample of $n = 10$ from the population:

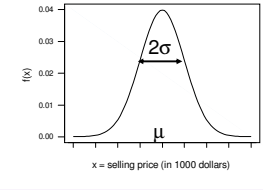
$$\{x_1, x_2, \dots, x_n\}$$

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Assumptions

[CI Estimation of mean– continued]

1. Assume that the random variable of interest X (selling price) is approximately normally distributed with mean μ and sd σ , or the sample



$\mu =$ unknown parameter of interest

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Ex 1:[CI Estimation of mean– continued]

$\mu =$ unknown population mean is estimated by the SAMPLE MEAN:

$$\hat{\mu} = \bar{x} = \frac{x_1 + x_2 + \dots + x_n}{n} = \frac{\sum_{i=1}^n x_i}{n}$$

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Ex 1:[CI Estimation of mean– continued]

A 95% confidence interval (CI) of the unknown mean is a random interval (L, U) such that

$$P(L < \mu < U) = 0.95$$

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Ex 1:[CI Estimation of mean– continued]

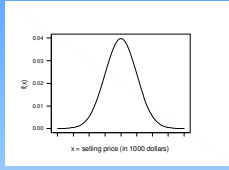
We will use the CENTRAL LIMIT THEOREM (CLT) : pdf of sample mean is approximately normal with mean μ and sd = σ/\sqrt{n} for large n . For many situations, $n \geq 25$ is sufficient.

$$\hat{\mu} = \bar{x} = \frac{x_1 + x_2 + \dots + x_n}{n} = \frac{\sum_{i=1}^n x_i}{n} \sim N\left(\mu, sd = \frac{\sigma}{\sqrt{n}}\right)$$

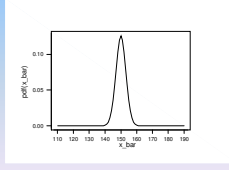
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CLT

pdf of selling price x



pdf of mean selling price \bar{x}



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Ex 1:[CI Estimation of mean– continued]

- 1 $\hat{\mu} = \bar{x} = \frac{x_1 + x_2 + \dots + x_n}{n} = \frac{\sum_{i=1}^n x_i}{n} \sim N(\mu, sd = \frac{\sigma}{\sqrt{n}})$
- 2 $z = \frac{\bar{x} - \mu}{\sigma / \sqrt{n}} \sim N(0,1)$

In words: (1) estimate of μ is the sample mean, which has the normal distribution with mean μ and standard deviation σ/\sqrt{n} ,
 (2) $(\bar{x} - \mu)$, standardized by its sd, has the standard normal distribution.

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Ex 1:[CI Estimation of mean– continued]

If you don't care for algebra,
please skip to slide 16.

To calculate 95% CI for mean μ , find z_0 from standard normal table so that $P(-z_0 \leq Z \leq z_0) = .95$.
 $-z_0 = -1.96, z_0 = +1.96$

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Ex 1:[CI Estimation of mean– continued]

$$z = \frac{\bar{x} - \mu}{\sigma / \sqrt{n}} \sim N(0,1)$$

$$P(-1.96 < \frac{\bar{x} - \mu}{\sigma / \sqrt{n}} < +1.96) = 0.95$$

Rewrite above as

$$P(\bar{x} - 1.96 \frac{\sigma}{\sqrt{n}} < \mu < \bar{x} + 1.96 \frac{\sigma}{\sqrt{n}}) = 0.95$$

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Ex 1:[CI Estimation of mean– continued]

95% CI for μ is

$$(\bar{x} - 1.96 \frac{\sigma}{\sqrt{n}}, \bar{x} + 1.96 \frac{\sigma}{\sqrt{n}})$$

Example calculation of CI:

Suppose the sample mean of 10 homes turned out to be \$147,000, and suppose the population sd $\sigma = \$10,000$

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Ex 1:[CI Estimation of mean– continued]

95% CI for μ is

$$(\bar{x} - 1.96 \frac{\sigma}{\sqrt{n}}, \bar{x} + 1.96 \frac{\sigma}{\sqrt{n}})$$

$$= (147 - \frac{1.96 \times 10}{\sqrt{10}}, 147 + \frac{1.96 \times 10}{\sqrt{10}})$$

$$= (147 - 6.20, 147 + 6.20)$$

$$= (140.8, 153.2)$$

$$= (\$140800, \$153200)$$

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Ex 1:[CI Estimation of mean– continued]

The above formula assumed that σ is known – which is rarely the case.

When σ is unknown – it is estimated by the sample sd s .
 Suppose for example 1 that sample sd of 10 observations was \$9,000.

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Ex 1:[CI Estimation of mean– continued]

$$\hat{\sigma} = s = \sqrt{\frac{\sum_{i=1}^n (x_i - \bar{x})^2}{n-1}}$$

$z = \frac{\bar{x} - \mu}{\sigma / \sqrt{n}}$ becomes

$t = \frac{\bar{x} - \mu}{s / \sqrt{n}} \sim t\text{-distribution with degrees of freedom } n-1$

In plain English, the distribution of $(\bar{x} - \mu)$ standardized by sample sd is a t-distribution with DEGREES OF FREEDOM $(n-1)$.

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Ex 1:[CI Estimation of mean– continued]

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Ex 1:[CI Estimation of mean– continued]

95% CI for μ becomes

$$\left(\bar{x} - 2.262 \frac{s}{\sqrt{n}}, \bar{x} + 2.262 \frac{s}{\sqrt{n}}\right)$$

which is sometimes expressed as

$$\bar{x} \pm 2.262 \frac{s}{\sqrt{n}}$$

Estimate \pm Reliability coefficient from t-table with $n-1$ degrees of freedom \times (sd of estimate)

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Ex 1:[CI Estimation of mean– continued]

Sample mean = \$147,000, sd = \$9,000

95% CI for μ is

$$\left(\bar{x} - 2.262 \frac{\sigma}{\sqrt{n}}, \bar{x} + 2.262 \frac{\sigma}{\sqrt{n}}\right)$$

$$= \left(147 - \frac{2.262 \times 9}{\sqrt{10}}, 147 + \frac{2.262 \times 9}{\sqrt{10}}\right)$$

$$= (147 - 6.44, 147 + 6.44)$$

$$= (140.6, 153.4)$$

$$= (\$140600, \$153400)$$

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CI Estimation of mean– continued]

95% confidence intervals in Ex.1 calculated by:

$$\left(\bar{x} - z_{0.975} \frac{\sigma}{\sqrt{n}}, \bar{x} + z_{0.975} \frac{\sigma}{\sqrt{n}}\right), \sigma \text{ known}$$

$z_{0.975} = 1.96$ from z-table

$$\left(\bar{x} - t_{df, 0.975} \frac{s}{\sqrt{n}}, \bar{x} + t_{df, 0.975} \frac{s}{\sqrt{n}}\right), \sigma \text{ unknown}$$

$t_{df, 0.975} = 2.262$ from t-table with $df = n-1$

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CI Estimation of mean– continued]

90% confidence intervals in Ex.1 will be calculated by:

$$\left(\bar{x} - z_{0.95} \frac{\sigma}{\sqrt{n}}, \bar{x} + z_{0.95} \frac{\sigma}{\sqrt{n}}\right), \sigma \text{ known}$$

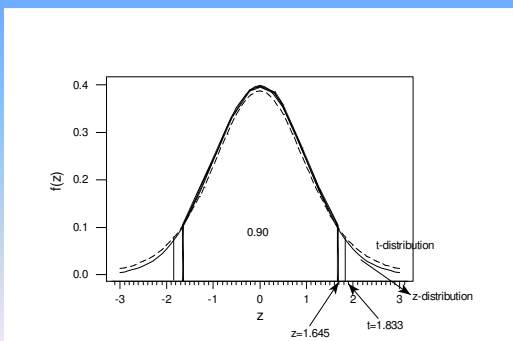
$z_{0.95} = 1.645$ from z-table

$$\left(\bar{x} - t_{df, 0.95} \frac{s}{\sqrt{n}}, \bar{x} + t_{df, 0.95} \frac{s}{\sqrt{n}}\right), \sigma \text{ unknown}$$

$t_{df, 0.95} = 1.833$ from t-table with $df = n-1=9$

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z- and t- table values for 90% confidence



REMARK:

For n larger than 200, the t-table and the z-table will give identical values:

$$t_{df, 1-\frac{\alpha}{2}} = z_{1-\frac{\alpha}{2}} \text{ for } n > 200$$

$100(1-\alpha) =$ confidence coefficient

$\alpha=0.05 \rightarrow 100(1-0.05) = 95\%$ confidence

$\alpha=0.10 \rightarrow 100(1-0.10) = 90\%$ confidence

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Example [CI Estimation of Binomial Proportion]

Suppose a random sample of $n=50$ inline skates are tested, and $x=3$ skates are found to be defective.

$X =$ #(defective skates) in sample of 50 has the BINOMIAL distribution

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Example [CI Estimation of Binomial Proportion – contd.]

$$X \sim \text{BIN}(n, p)$$

$n=50 =$ # of samples

$x =$ # of defectives in the sample

$p =$ proportion of defective skates in the population

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Example [CI Estimation of Binomial Proportion – contd.]

The unknown proportion of defectives p is estimated by the sample proportion of defectives:

$$\hat{p} = \frac{\text{\#(defectives) in sample}}{\text{\# of samples tested}} = \frac{x}{n}$$

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Example [CI Estimation of Binomial Proportion – contd.]

Following result follows from CLT:

$$\hat{p} \sim N\left(p, sd = \sqrt{\frac{p(1-p)}{n}}\right)$$

for n large so that $n\hat{p} \geq 5$, $n(1-\hat{p}) \geq 5$

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Example [CI Estimation of
Binomial Proportion – contd.]

$$sd(\hat{p}) = \sqrt{\frac{p(1-p)}{n}} \text{ is estimated by}$$

$$sd(\hat{p}) = \sqrt{\frac{\hat{p}(1-\hat{p})}{n}}$$

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Example [CI Estimation of
Binomial Proportion – contd.]

Using the formula for CI (see slide 21)
estimate \pm reliability coefficient \times sd(estimate)
we get the 95% CI for p as:

$$\hat{p} \pm 1.96 \times \sqrt{\frac{\hat{p}(1-\hat{p})}{n}}$$

The reliability coefficient of 1.96 is
from the z-table or the standard
normal table.

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Example [CI Estimation of
Binomial Proportion – contd.]

In the example, $x = 6$, $n = 50$, so

$$\hat{p} = \frac{x}{n} = \frac{6}{50} = 0.12$$

$$sd(\hat{p}) = \sqrt{\frac{\hat{p}(1-\hat{p})}{n}}$$

$$= \sqrt{\frac{0.12(1-0.12)}{50}} = \sqrt{.002112} = .045957$$

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Example [CI Estimation of
Binomial Proportion – contd.]

95% CI:

$$\hat{p} \pm 1.96 \times sd(\hat{p})$$

$$= 0.12 \pm 1.96 \times .045957$$

$$= 0.12 \pm 0.09$$

$$= (0.03, 0.21)$$

In words: we
are 95%
confident that
p lies
between 3%
and 21%.

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